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# The Relationship Between Defence Expenditure, Trade Openness and Foreign Direct Investments: The Case of BRICS and MINT Countries

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#### Abstract

This study examines the relationship among trade openness, foreign direct investments and defence expenditures for the BRICS and MINT countries in the 1990-2019 period. As analysis method, it was estimated by Pesaran (2008) Cross-Section Dependency test, Hadri-Kuruzomi (2012) panel unit root test, Pesaran (2006) CCE test and Emirmahmutoglu and Kose (2011) panel causality tests. According to the test findings, a bidirectional causality relationship was found between defense expenditures and foreign direct investments for BRICS countries. In addition, for MINT countries, there is a one-way causality relationship from defense expenditures to foreign direct investments and from trade openness to defense expenditures.

**Keywords:** Defence Expenditures, Trade Openness, Foreign Direct Investment, Panel Causality Analysis. **DOI:** 10.7176/JESD/13-4-01

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#### 1. Introduction

Developing countries prefer foreign resources, foreign borrowing or foreign direct investments as a source of financing while realizing their economic development goals (Cestepe et al., 2013: 2). Foreign direct investments, which are important for developing countries, create positive externalities by providing technology and knowledge transfer as well as being a source of finance. In addition, these investments can increase the integration of countries with the global economy (Angelopolou and Liargovas, 2014: 471).

It has been observed that foreign direct investments as a source of foreign financing have increased in the economic development processes of underdeveloped countries since the 1990s. Foreign direct investments are among the important actors of globalization (Yilmazer, 2010: 242, 243). As Chakrabarti (2001), trade openness is one of the determinants of foreign direct investment. According to Torissi (1985), having a trade surplus also positively affects foreign direct investments in a well-functioning economy (Uzun, 2010: 111).

The determinants factors of foreign direct investments include market volume, economic stability, trade openness, exchange rate, capital stock, infrastructure (share of public investments in the budget in energy and transportation), political risk, structural reforms, human capital (Yılgor et al., 2011).

In the economics literature, the relationship between defence expenditures, foreign direct investments and trade openness variables with growth has been examined more. In addition, the relationship between foreign direct investments and defence expenditures, foreign direct investments and trade openness or the relationship between defence expenditures and trade openness has been studied in the literature. As a result of detailed literature review, no study has been found that handles these three variables together.

Although there have been many studies in the literature on defence spending, foreign direct investment and trade openness, the rising powers of the New World Order (NWO), BRICS (Brazil, Russia, India, China, South Africa) and MINT (Mexico, Indonesia, Nigeria and Turkey), the study is intended to contribute to the literature in terms of the limited number of studies examining this country groups.

Due to the increase in globalization and financial liberalization tendencies that led to an increase in capital movements in the world in the 1980s, the relationship between defense expenditures, foreign direct investments, trade openness for the period 1990-2019 was analyzed by the panel causality test by Emirmahmutoglu and Kose (2011). Thus, the analysis findings can be analyzed both in general and for each country.

#### 1. The Relationship Between Trade Openness and Foreign Direct Investment

Trade openness is the determinant of foreign direct investment inflows. Foreign direct investments are important not only for exports but also for imports of intermediate and capital goods. In every respect, foreign direct investments increase the trade volume. Trade openness is also a determinant of foreign direct investment. Nunes et al. (2006) and Sahoo (2006), the ratio of exports and imports to GDP expresses trade openness (Vijayakumar et al., 2010: 5, 6).

Country's trade openness is calculated by the total export and import/GDP ratio (Keskin, 2020: 19). The concept of trade openness is aimed to create an international free trade environment by removing state controls on the import and export of goods and services. It is explained as the elimination of restrictions that prevent the competitive environment by ensuring the free movement of labor and capital, as well as goods and services, between countries (Yaprakli, 2007: 68).

Trade openness is one of the effective arguments for attracting foreign direct investment to a country. While foreign direct investments prefer the country, countries' exchange rate, market volume, trade and financial openness, political stability, labor and investment costs, current account deficit, foreign debt, budget deficit, inflation, energy use level, human capital, tax rates, public expenditures. It also takes into account the macroeconomic and political stability of the countries. As Chakrabarti (2001), the degree of openness of countries to foreign trade is a determining and effective factor in foreign direct investment inflows (Karis and Ayla, 2018: 256, 257).

Whether it prefers countries with stable foreign direct investments or whether it stabilizes the country's economies is among the controversial issues in the literature (Okuyan and Erbaykal, 2007: 48). In a country with high openness in trade, the formation of an environment of extreme competition and the hypothesis of "tariff jumping - trade restrictions" can negatively affect the economies of the countries (Keskin, 2020: 19, 20).

To investigate the relationship between trade openness and foreign direct investment, Edwards (1990), Gastanaga et al. (1998), Chakrabati (2001), Asiedu (2001), Kandiero and Chitiga (2006), Yaprakli (2006), Demirhan and Masca (2008), Snow and Sweet Words (2008), Tari and Bidirdi (2009), Jadhav (2012) ). ), Liargovas and Skandalis (2012), Cantah et al. (2013), Asghar (2016), Dima (2016), Tampakoudis et al. (2017), Terzi and Bekar (2017) concluded that there is a positive relationship between trade openness and foreign direct investment, and that trade openness is determinant in foreign direct investment inflows.

In the study of Ozcan and Arı (2010), a negative relationship was found between the two variables. Schmitz and Bieri (1972), Busse and Hefeker (2007), Vijayakumar et al. (2010), Miškinis and Juozenaite (2015), Tsaurai (2015), and Belloumi (2014) found a neutral relationship, that is, a statistically significant relationship.

#### 2. The Relationship Between Defence Expenditures and Foreign Direct Investments

In the literature, the relationship of defense expenditures with growth and foreign trade has been examined, but the relationship between defense expenditures and foreign direct investment, which is seen as the driving factor of growth, has not been analyzed much (Pacific, Shan and Ramazan). , 2017: 577, 578). It is a tool in that defense expenditures are a variable related to GDP per capita and affect foreign direct investments over GDP per capita. However, defense expenditures do not directly affect foreign investment (Hussain and Kimuli, 2012: 20).

With regard to post-Cold War military spending, Drezner and Hite-Rubin (2016) state that if countries spend and invest more in their armies, foreign direct investment will lead to safer property rights and lead to more investment. Norrlof (2010) explains that military power will affect private sector investment decisions and create a safer investment environment with the concept of "Geoeconomic Nepotism". According to the geoeconomic favoritism hypothesis, it is explained that the increase in military power due to the increase in the defense expenditures of the countries will positively affect the foreign direct investments as it ensures the security of property rights (Drezner and Hite-Rubin, 2016: 2).

When a detailed literature review is made, in the studies of Kennedy (1974), Benoit (1978), Whynes (1979), Barro and Sala-i-Martin (2004), Dunne, Smith and Willenbockel (2005), defense expenditures affect the trade and economic trust environment of countries. They say it will increase economic growth by raising it.

Rothschild (1973), Smith (1980), Deger and Smith (1983), Leontief and Duchin (1983), Lim (1983), Landau (1985), Mintz and Huang (1990), Ram (1995), Dunne, Nikolaidou and In Smith (2002)'s studies, it is claimed that defense expenditures exclude productive sector investments and reduce economic growth (Aziz and Asadullah, 2017: 1; Aziz and Khalid, 2019: 240). In the studies of Acemoglu and Robinson (2012), it is argued that defense expenditures - the crowding out effect - will exclude foreign direct investments (Drezner and Hite-Rubin, 2016).

#### 3. The Relationship Between Trade Openness and Defence Spending

Although there are studies on the macroeconomic determinants of defense expenditures, it is stated by Acemoglu and Yared (2010) that defense expenditures are associated with trade openness as an indicator of globalization. For this reason, the concept of trade openness is discussed in the literature mostly with the globalization dimension.

In the literature, defense expenditures are mostly examined in relation to economic growth. The relationship between globalization and defense expenditures was first reported by Dreher et al. (2006) examined and it was concluded that globalization has no effect on defense expenditures. Analysis findings of Nikolaiodu's (2008) study were also reported by Dreher et al. (2006) in the same direction.

Dunne et al. (2008) and Seiglie (2016) concluded that trade openness has a positive effect on defense spending. In Irondoust's (2007) study, it was concluded that defense expenditures increased with globalization. Solarin (2017) concluded that globalization both increased the military burden of countries and reduced their defense expenditures.

In the study of Kurt and Kilic (2019), it was concluded that economic, political and social globalization increased defense expenditures. Acemoglu and Yared (2010) concluded that the increase in defense expenditures causes a decrease in the terms of trade of the countries. In the studies of Cengiz and Manga (2020), it was determined that political globalization increased defense expenditures, economic globalization decreased defense expenditures and social globalization had no effect on defense expenditures.

There are various opinions in the literature about the relationship between globalization and defense expenditures. In addition to the views that globalization has increased defense expenditures through social expenditures, there are also opinions that globalization has negative effects on defense expenditures and that defense expenditures should be reduced with the effect of globalization and economic liberalization. As Paul and Ripsman (2004) stated, trade and economic liberalization and the pursuit of increasing wealth force countries to use military tools. Free trade regulations also make it possible to increase defense expenditures (Solarin, 2017: 854, 855).

### 4. Data and Methodology

In this study, the causality relationship between defense expenditures, openness to trade and foreign direct investments for the period 1990-2019 and for BRICS and MINT countries is analyzed with the Emirmahmutoglu-Kose (2011) panel causality test. The variables used in the analysis, defense expenditures and foreign direct investment data were taken from 'data.worldbank.org' database and trade openness data were taken from 'theglobaleconmy.com'. The application is made using Gauss 10 and Stata 12 econometrics programs.

Panel causality test, which is one of the second generation tests, was applied for the group of emerging market economies, BRICS and MINT countries, in terms of the relationship between defense expenditures, openness to trade and foreign direct investments. The relationship between the variables was analyzed by Pesaran (2008) Cross-Section Dependency test, Hadri-Kuruzomi (2012) panel unit root test and long-term regression coefficients of the variables Pesaran (2006) CCE (Common Correlation Effect) test and Emirmahmutoglu and Kose (2011) panel causality tests has been analyzed.

The econometric models of the study are analyzed for 3 different models as stated 1, 2 and 3 below:

Trade =  $a + \beta 1Defence + \beta 2Fdi + u$ 

$$Defence = a + \beta 1 Trade + \beta 2 Fdi + u$$
(2)  
FDI = a + \beta 1 Defence + \beta 2 Trade + u (3)

In order to determine the analysis methods, firstly, the analysis of cross-section dependence and heterogeneity of the series, cointegration and causality analyzes were applied to the series after the unit root.

## 4.1.1. Pesaran CD (2004) Cross Section Dependency Test

With the cross-sectional dependency test, it is to examine whether the units are dependent on each other, in other words, whether a shock to a series and all cross-section units has the same effect from this shock to the series. Cross-section dependency was investigated using the Pesaran CDLM (2004) test. CDLM test statistic was derived with this cross section dependency test and this test was chosen because it can be used in cases where both N> T and T> N. (Ozturk, 2018: 4, 5):

$$CD = \sqrt{\frac{2T}{N(N-1)}} \sum_{i=1}^{N-1} \sum_{j=i+1}^{N} T \tilde{\rho}_{ij}^2 - 1 \sim N(0,1)$$
(1)

The cross section dependency test is examined with CDLM test Ho zero hypothesis:  $H_0: \rho_{ij} = \rho_{ij} = cor(u_{it}, u_{jt}) = 0, i \neq j.$ 

## 4.1.2. Pesaran and Yamagata (2008) Homogeneity Test

The homogeneity test for the cross sections of the slope parameters for the cointegration model was performed by Swamy (1970). Pesaran and Yamagata (2008) test N and T have different sizes and make it possible to test the homogeneity assumption. Developing two different homogeneity tests with the Swamy test, the  $\beta i$  coefficients in cointegration models are analyzed with equation 2 and panel regression equation (Ozturk, 2018: 5):

$$Y_{it} = \alpha + \beta_i X_{it} + \varepsilon_{it}$$
(2)  
Using the equation 2, the delta tilde value is calculated for small samples,  
$$\chi = \sqrt{N} \frac{N^{-1} \breve{S} - K}{N^{-1} \breve{S} - K}$$
(2)

$$\tilde{\Delta} = \sqrt{N} \left( \frac{N - S - K}{\sqrt{2K}} \right) \sim X_k^2 \tag{3}$$

Equation 4, the corrected form of the delta test statistic gives the delta tilda adj. value:

(4)

$$\tilde{\Delta}_{adj} = \sqrt{N} \left( \frac{N^{-1} \check{S} - k}{v(T, k)} \right) \sim N(0, 1)$$

For the delta test, Ho: slope coefficient is homogeneous and H1: slope coefficient is interpreted as not homogeneous ie heterogeneous (Pesaran and Yamagata, 2008: 56).

BRICS Countries						
Variables	Defence		Trade		Fdi	
CD Tests	Stat.	Prob.	Stat.	Prob.	Stat.	Prob.
Cd Lm2 (Pesaran 2004)	32.350**	0.000	15.081**	0.000	4.547**	0.000
Cd LM (Pesaran 2004 CD)	-3.423**	0.000	-3.672**	0.000	-3.226**	0.001
Bias-adjusted CD test	15.212**	0.000	9.996**	0.000	1.321*	0.093
Delta Homogenity Test Res	sults for Mod	lels				•
Homogenity Test	Mod	el 1	Mod	el 2	Mod	el 3
Delta	52.57		39.25		68.30	
	(0.000)		(0.000)		(0.000)	
Delta-tilda-adj	54.49		38.96		13.35	
-	(0.000)		(0.000)		(0.000)	
		MINT Cou	intries			
Variable	Defe	nce	Trade		Fdi	
CD Tests	Stat.	Prob.	Stat.	Prob.	Stat.	Prob.
Cd Lm2 (Pesaran 2004)	16.474**	0.000	21.032**	0.000	16.104**	0.000
Cd I M (Pesaran 2004 CD)	1 1 0 0 1 1					
Cu Livi (i csarali 2004 CD)	-4.189**	0.000	-3.912**	0.000	-4.244**	0.000
Bias-adjusted CD test	-4.189** 10.399**	0.000	-3.912** 3.360**	0.000 0.000	-4.244** 1.335*	0.000 0.091
Bias-adjusted CD test Delta Homogenity Test Res	-4.189** 10.399** sults for Mod	0.000 0.000 lels	-3.912** 3.360**	0.000	-4.244** 1.335*	0.000 0.091
Bias-adjusted CD test Delta Homogenity Test Res Homogenity Test	-4.189** 10.399** sults for Mod Mod	0.000 0.000 lels el 1	-3.912** 3.360** Mod	0.000 0.000 el 2	-4.244** 1.335* Mod	0.000 0.091 el 3
Bias-adjusted CD test Delta Homogenity Test Res Delta	-4.189** 10.399** sults for Mod 9.01	0.000 0.000 lels el 1	-3.912** 3.360** Mod 54.37	0.000 0.000 el 2	-4.244** 1.335* Mod 46.02	0.000 0.091 el 3
Bias-adjusted CD test Delta Homogenity Test Res Homogenity Test Delta	-4.189** 10.399** sults for Mod 9.01 (0.001)	0.000 0.000 lels el 1	-3.912** 3.360** <b>Mod</b> 54.37 (0.006)	0.000 0.000 el 2	-4.244** 1.335* Mod 46.02 (0.001)	0.000 0.091 el 3
Bias-adjusted CD test Delta Homogenity Test Homogenity Test Delta Delta Delta-tilda-adj	-4.189** 10.399** sults for Mod 9.01 (0.001) 38.5	0.000 0.000 lels el 1	-3.912** 3.360** <b>Mod</b> 54.37 (0.006) 10.67	0.000 0.000 el 2	-4.244** 1.335* Mod 46.02 (0.001) 24.86	0.000 0.091 el 3

Table 1. Cross Section Dependency and Homogenity Tests Results

Note: \*\*\*, \*\*, \* indicate 10%, 5% and 1% significance levels, respectively Trade: Trade Openness, Defence: Defence Expenditure, Fdi: Foreing Direct Investment.

The zero (Ho) hypothesis was rejected because the probability values were less than 0.05 of the defence expenditures, trade openness and foreign direct investment variables calculated from table 1 so it was concluded that there is a cross-sectional dependency in series. There is a cross-sectional dependency between countries belonging to each country group that constitutes the panel. In other words, the defence spending, trade openness and foreign direct investment shock coming to one of the countries affect other country groups as well.

Since the probability values of delta test statistics are less than 0.05, the zero hypothesis is rejected. In other words, it is concluded that the constant term and slope coefficients are not homogeneous in the cointegration equation. Therefore, from the homogeneity test results, it is decided that the cointegration test interpretations of the countries in the panel are valid and reliable (Pesaran and Yamagata, 2008).

4.1.3. Hadri-Kurozumi (2012) Panel Unit Root Test

Hadri-Kurozumi (2012) panel unit root test is the adapted form of Kwiatkowski, Phillips, Schmidt and Shin (KPSS, 1992) test for panel data analysis. Unit root test was developed considering the cross-sectional dependency. This test is assumed to have a normal distribution (Temiz and Konat, 2019: 2335):

$$Y_{it} = z_t' \delta_i + f_t \gamma_i + \varepsilon_{it}$$

(5)

Using Equation 5, the presence of the unit expressed in root panel unit root model being tested.

Table 2. Hauri-Kurozann (2012) Onit Root Test Result						
BRICS Countries						
Variables	Con	stant	<b>Constant and Trend</b>			
	Z <sup>SPAC</sup> (prob)	Z <sup>LA</sup> (prob)	Z <sup>SPAC</sup> (prob)	Z <sub>A</sub> <sup>LA</sup> (prob)		
Defence	-1.6305	1.6116	0.2353	0.4513		
	(0.5015)***	(0.5035)***	(0.4070)***	(0.3259)***		
Trade	-1.0262	-1.0474	-1.4649	-1.4208		
	(0.8476)***	(0.8525)***	(0.9285)***	(0.9223)***		
Fdi	-0.4810	-0.3877	-0.5836	-0.3324		
	(0.6847)***	(0.6509)***	(0.7203)***	(0.6302)***		
		MINT Count	ries			
	Cons	tant	Constant	and Trend		
Variables	Z <sub>A</sub> <sup>SPAC</sup> (prob)	Z <sup>LA</sup> (prob)	Z <sub>A</sub> <sup>SPAC</sup> (prob)	Z <sub>A</sub> <sup>LA</sup> (prob)		
Defence	13.4247	14.683	39.103	46.500		
	(0.5775)***	(0.1118)***	(0.8396)***	(0.8793)***		
Trade	0.8192	1.1411	3.7938	4.3045		
	(0.2063)***	(0.1269)***	(0.7092)***	(0.4087)***		
Fdi	-1.0812	-0.8684	0.8306	1.5689		
	(0.8602)***	(0.8074)***	(0.2031)***	(0.5083)***		

Table 2. Hadri-Kurozumi (2012) Unit Root Test Result

Note: The optimum lag lengths for the variables shown in the "L" column in the table were determined with the Schwarz information criterion. The (\*\*\*) sign in front of the test statistics obtained from the constant + trend forms indicates that the basic hypotheses for the variables are accepted at the 1% significance level. Respectively Trade: Trade Openness, Defence: Defence spending, Fdi: Foreing Direct Investment.

Hadri-Kurozumi (2012) panel unit root test for the variables of defence expenditures, trade openness and foreign direct investments of both BRICS and MINT countries, when the values with constant and constant and trend are interpreted in accordance with the analysis findings given in table 2, it is determined for both selected country groups. Since the null hypothesis cannot be rejected at the 1% significance level, it is concluded that the variables are stationary.

## 4.1.4. Estimating Long Term Cointegration Coefficients

Pesaran (2006) gives better results with the CCE estimator in small samples and analysis methods that do not consider cross-sectional dependence (Nazlıoğlu, 2010: 102). With the Pesaran (2006) CCE (Common Correlation Effects) estimator, the long-term cointegration coefficients of the variables are tested by considering the heterogeneity and cross-sectional dependence, and the regression coefficients are estimated for each crosssectional unit. (Keskin and Aksoy, 2019: 9). In case of N > T and T > N, the CCE model can be used. In the CCE model, the long-term regression cointegration coefficients are estimated by two separate estimators, CCEMG (Commonly Associated Effects Average Estimator) and CCEP, given in equations 6 and 7. (Usually Associated Effects are Combined) (Pesaran, 2008: 52):

$$\hat{b}_{CCEMG} = \frac{1}{N} \sum_{i=1}^{N} \hat{b}_i$$

$$b_{CCEP}' = \left(\sum_{i=1}^{N} \theta_i x_i M_w x_i\right)^{-1} \left(\sum_{i=1}^{N} \theta_i x_i M_w x_i\right)$$
(6)
(7)

**TII 2 D** 

1 abio 3. Pesaran (2006) CCE Estimation Test Result						
BRICS Countries						
Models	Variables	Coef.	t-stat			
Model 1	Defence	1.4391	0.25			
Dependent Variable: Trade	Fdi	2.4391	1.47			
	Wald chi2 : 2.27 pro	b > chi2 : 0.322				
Model 2	Trade	-0.0038	-0.45			
Dependent Variable: Defence	Fdi	-0.0436	-1.66			
Expenditures	Wald chi2 : 5.22 pro	bb > chi2 : 0.073 **				
Model3	Defence	-1.5563	-1.54			
Dependent Varible: Fdi	Trade	0.06209	2.16			
	Wald chi2 : 4.73 pro	b > chi2 : 0.093 **	•			
	MINT Countries					
Models	Variables	Coef.	t-stat			
Model 1	Defence	24.788	1.29			
Dependent Variable: Trade	Fdi	-1.2652	-0.60			
	Wald chi2 : 3.02 pro	b > chi2 : 0.221				
Model 2	Trade	0.0014	0.92			
Dependent Variable: Defence	Fdi	0.0720	2.44			
Expenditures	Wald chi2 : 7.48 pro	b > chi2 : 0.023 **				
Model 3	Defence	1.599	1.27			
Dependent Variable: Fdi	Trade	-0.002	-0.01			
	Wald chi2 : 1.76 pro	b > chi2 : 0.414				

(AAAA) COEE .

Note: \*\*\*, \*\*, \* indicate 10%, 5% and 1% significance levels, respectively Trade: Trade Openness, Defence: Defence Expenditure, Fdi: Foreing Direct Investment.

When the results expressed in table 3 are interpreted, the long-term coefficients can be interpreted since models 2 and 3 for BRICS countries and model 2 for MINT countries (probe values<0.05) are significant. According to model 2, where defence expenditures are taken as dependent variables for BRICS countries, it is concluded that a 1% increase in trade openness reduces defence expenditures by 0.0038%, while a 1% increase in foreign direct investments reduces defence expenditures by 0.0436%.

According to model 3, where the foreign direct investment variable is taken as the dependent variable, it is concluded that in BRICS countries, a 1% increase in defence expenditures decreases foreign direct investments by 1.556% and a 1% increase in trade openness increases foreign direct investments by 0.062%.

For MINT countries, according to model 2, where defence expenditures are taken as the dependent variable, it is concluded that a 1% increase in trade openness increases defence expenditures by 0.0014% and a 1% increase in foreign direct investments increases defence expenditures by 0.0720%.

#### 4.1.5. Emirmahmutoglu ve Kose (2011) Panel Causality Test

Since Emirmahmutoglu and Kose (2011) is a test based on the Toda-Yamamoto (1995) causality test in the panel causality test time series, the standard panel VAR estimation is first made and the appropriate lag length (p) is determined. In the next step, the integration level (dmax) of the variable with the highest integration degree is added to the p lag. The level values of the series for the (p + dmax) delay are based on estimating with the panel VAR model (Sahin and Durmus, 2019: 195). It can be applied in case of cross section dependency. It is possible to analyze by using different level values of the series as I(0) and I(1). As expressed in Equations 8 and 9, a causality analysis is applied for two variables by establishing the VAR model (Ilgaz Yildirim and Sahin, 2018: 449):

$$x_{i,t} = \mu_i^x + \sum_{\substack{j=1\\k_i+dmax_i}}^{k_i+dmax_i} A_{11,ij} x_{i,t-j} + \sum_{\substack{j=1\\k_i+dmax_i}}^{k_i+dmax_i} A_{12,ij} x_{i,t-j} + \mu_{i,t}^x$$
(8)

$$x_{i,t} = \mu_i^x + \sum_{j=1}^{M_i + m_i} A_{21,ij} x_{i,t-j} + \sum_{j=1}^{M_i + m_i} A_{22,ij} x_{i,t-j} + \mu_{i,t}^y$$
(9)

 $d_{max}$  expresses the maximum degree of integration for each i. In line with the estimation results, a modified Wald (MWALD) test is applied for the  $k_i$  delay. According to Equation 8, the  $H_0$  hypothesis of the test is that there is a causality relationship from Y to X.

Causality Direction	Panel Fisher	P-val			
BRICS Countries					
Fdi to Trade	4.927	0.896			
Trade to Fdi	3.004	0.981			
Defence Expenditures to Fdi	19.476	0.035**			
Fdi to Defence Expenditues	18.252	0.051**			
Trade to Defence Expenditures	6.314	0.788			
Defence Expenditures to Trade	7.265	0.700			
MINT	Countries				
Fdi to Trade	9.257	0.321			
Trade to Fdi	2.407	0.966			
Defence Expenditure to Fdi	15.679	0.047**			
Fdi to Defence Expenditures	11.761	0.162			
Trade to Defence Expenditures	17.909	0.022**			
Defence Expenditures to Trade	10.593	0.226			

## Table 4: Emirmahmutoglu ve Kose (2011) Panel Causality Direction Result

Note: \*\*\*, \*\*, \* indicate 10%, 5% and 1% significance levels, respectively Trade: Trade Openness, Defence: Defence Expenditures, Fdi: Foreing Direct Investment.

According to panel causality analysis in panel general findings in table 4, for BRICS countries it has been found that there is a unidirectional causality relationship between defence expenditures and foreign direct investments, while a unidirectional causality relationship from defence expenditures to foreign direct investments and from trade openness to defence expenditures in MINT countries it has been concluded.

Fdi to Trade				Tr	ade to Fdi		
BRICS Countries							
i	Lag	Wald	p-val	Lag	Wald	p-val	
Brazil	1.000	0.005	0.945	1.000	0.009	0.925	
Russia	1.000	0.020	0.887	1.000	0.349	0.554	
China	1.000	0.186	0.666	1.000	0.348	0.555	
India	1.000	1.180	0.178	1.000	0.065	0.798	
South Africa	1.000	0.034	0.854	1.000	0.001	0.979	
Panel Fisher : 4.927				Panel Fishe	er : 3.004		
p-value : 0.896				p-value	: 0.981		
		MI	NT Countries				
i	Lag	Wald	p-val	Lag	Wald	p-val	
Mexican	1.000	1.908	0.167	1.000	0.031	0.859	
Indonesia	1.000	0.030	0.862	1.000	0.550	0.458	
Nigeria	1.000	0.036	0.850	1.000	0.026	0.872	
Turkey	1.000	3.069	0.080**	2.000	3.123	0.873	
Panel Fisher : 9.257				Panel Fisher : 2.407			
p-value : 0.321				p-value	: 0.966		

 Table 5. Emirmahmutoglu – Kose (2012) Panel Causality Test Results For Countries

Note: \*\*\*, \*\*, \* indicate 10%, 5% and 1% significance levels, respectively Trade: Trade Openness, Defence: Defence Expenditures, Fdi: Foreing Direct Investment.

In table 5, there is found that a causality relation from foreign direct investments to trade openness (10%) for Turkey in the group of MINT countries.

Defence to Fdi				Fdi to Defence			
BRICS Countries							
i	Lag	Wald	p-val	Lag	Wald	p-val	
Brazil	1.000	17.908	0.001**	4.000	7.590	0.108	
Russia	1.000	0.718	0.397	1.000	0.039	0.844	
China	1.000	0.132	0.716	1.000	0.430	0.512	
India	1.000	1.492	0.222	2.000	3.664	0.160	
South Africa	1.000	0.122	0.727	2.000	8.455	0.015**	
Panel Fisher: 19.476				Panel Fisher : 18.252			
p-value : 0.03	35**			p-value : 0.051**			
			MINT Countr	ies			
i	Lag	Wald	p-val	Lag	Wald	p-val	
Mexican	1.000	0.778	0.378	1.000	1.609	0.205	
Indonesia	1.000	6.126	0.013**	1.000	3.073	0.080**	
Nigeria	1.000	0.401	0.527	1.000	0.568	0.451	
Turkey	1.000	2.087	0.149	1.000	0.770	0.380	
Panel Fisher: 15.679			Panel Fisher: 11.761				
p-value : 0.047**			p-value : 0.162				

## Table 6. Emirmahmutoglu – Kose (2012) Panel Causality Test Results For Countries

Note: \*\*\*, \*\*, \* indicate 10%, 5% and 1% significance levels, respectively Trade: Trade Openness, Defence: Defence Expenditures, Fdi: Foreing Direct Investment.

In table 6, it is seen that the causality relation from defence expenditures to foreign direct investments is valid for Brazil in the group of BRICS countries and Indonesia in the group of MINT countries (5%). The causality relation from foreign direct investments to defence expenditures is valid for the BRICS countries in South Africa and for Indonesia in the MINT countries group.

Table 7. Emirmahmutoglu – F	Kose (2012)	) Panel Causality	Test Results F	or Countries
rable /. Enni manmatogia i	TOSC (TOTA)	f I and Causanty	I Cot Itcoulto I	or countries

Trade to Defence				Defence to Trade			
BRICS Countries							
i	Lag	Wald	p-val	Lag	Wald	p-val	
Brazil	1.000	0.008	0.927	1.000	0.190	0.663	
Russia	1.000	0.144	0.704	1.000	0.181	0.671	
China	1.000	2.925	0.087***	1.000	1.014	0.314	
India	1.000	0.073	0.788	1.000	0.312	0.577	
South Africa	1.000	0.004	0.948	1.000	0.956	0.328	
Panel Fisher : 6.314				Panel Fisher: 7.265			
p-value : 0.78	8			p-value : 0.700			
			MINT Countr	ies			
i	Lag	Wald	p-val	Lag	Wald	p-val	
Mexican	1.000	3.470	0.063**	1.000	0.862	0.353	
Indonesia	1.000	4.127	0.042**	1.000	1.442	0.230	
Nigeria	1.000	2.691	0.101	1.000	1.608	0.205	
Turkey	1.000	0.487	0.485	1.000	1.067	0.032**	
Panel Fisher : 17.909			Panel Fisher: 10.593				
p-value : 0.022**			p-value : 0.226				

Note: \*\*\*, \*\*, \* indicate 10%, 5% and 1% significance levels, respectively Trade: Trade Openness, Defence: Defence Expenditures, Fdi: Foreing Direct Investment.

In table 7, it is seen that the causality relation from trade openness to defence expenditures is found for China (10%) in the group of BRICS countries and Mexico (10%) and Indonesia (5%) in the group of MINT countries. The causality relationship from defence expenditures to trade openness is valid only for Turkey (5%) in the group of MINT countries.

## CONCLUSION

With the collapse of the Bretton-Woods system in the late 1970s and the conversion of different currencies to convertibility, the international rate of movement of capital increased. Financial institutions, whose profit rates

decreased in developed countries, turned to developing countries that offer higher returns through foreign direct investment in order to increase their profit rates.

This has accelerated the globalization process. The concept of globalization, which came to the fore with the changes in information, communication and technology after the 1980s, is a multidimensional concept due to its socio-economic, political and military aspects. Trade openness is an important indicator of globalization. As a result of liberalization of trade liberalization, financial liberalization and liberalization of capital movements with the effect of globalization, foreign direct investment inflows are seen as a source of technological innovation and employment as international capital movements.

In this study, the relationship between defense expenditures, foreign direct investments and trade openness for the BRICS and MINT group countries for the period 1990-2019 was examined. The analysis of the study was analyzed with Pesaran (2008) Cross-Section Dependency test, Hadri-Kuruzomi (2012) panel unit root test, and the long-term regression coefficients of the variables were determined by Pesaran (2006) CCE (Common Correlative) effect test and Emirmahmutoglu and Kose (2011) panel causality tests. has been estimated.

According to Hadri-Kurozumi (2012) panel unit root test, it was concluded that defense expenditures, trade openness and foreign direct investment variables in BRICS and MINT countries are stationary at 1% significance level according to fixed, constant and trend values and the null hypothesis cannot be rejected. In addition, according to the unit root test findings, since the variables are not stationary at the level, it can be interpreted as the convergence of defense expenditures, trade openings and foreign direct investment for BRICS and MINT country groups.

According to the Pesaran (2006) CCE estimator for the long-term cointegration coefficient result, a 1% increase in trade openness in BRICS countries reduces defense expenditures by 0.0038%, which is in line with the analysis findings of Solarin's (2017) study. It was concluded that a 1% increase in foreign direct investments decreased defense expenditures by 0.0436%, a 1% increase in defense expenditures decreased foreign direct investments by 1.556%, and a 1% increase in trade openness increased foreign direct investments by 0.062%. It was concluded that in MINT countries, a 1% increase in trade openness increased defense spending by 0.0014%, which is consistent with the analysis findings of Dunne et al. (2008), Seiglie (2016), Irondoust (2017) and Kurt and Kilic (2019) ' results show that a 1% increase in foreign direct investments in MINT countries increases defense expenditures by 0.0720%. According to Emirmahmutoğlu – Köse (2012) panel causality analysis results, for BRICS countries, there is a bidirectional causality relationship defence expenditures and foreign direct investments. In MINT countries has been concluded that unidirectional causality relation from defence expenditures to foreign direct investments and from trade openness to defence expenditures. That is, a one way causality relationship from defence expenditures to foreign direct investments has been found for both BRICS and MINT countries.

According to the findings of causality analysis, the existence of a one-way causality relationship from defence expenditures to foreign direct investments is interpreted as that the 'Geo-Economic Nepotism' hypothesis is valid for BRICS and MINT countries and that the hypothesis is valid for Brazil (BRICS) and Indonesia (MINT) countries in particular.

When the causality analysis on the basis of countries interprets the findings for Turkey, there is a causality relationship from foreign direct investments to trade openness and defence spending to trade openness is valid. The causality relationship from foreign direct investments to defence expenditures is valid for South Africa and Indonesia. The causality relationship from trade openness to defence expenditure is valid for China, Mexico and Indonesia.

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